

Urban Renewal, Gentrification, and Inequality

Evidence from Chicago Public Housing Demolitions

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Introduction

Urban renewal, segregation, and inequality

Place-based policies as a response to **urban inequality**

- Annual spending of \$100B by US governments to stimulate redevelopment (Story 2012; Kline and Moretti 2014)

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Policies focused on **urban renewal** have distinct features

- Destruction of housing occupied by lower-income residents
- Forced displacement of residents (especially racial minorities) [Map](#)

How can urban renewal programs change neighborhoods and cities?

Direct effects:

- Destruction of low-income housing
 - Preferences for the presence/absence of low-income housing: Diamond & McQuade (2019)
- Direct displacement of residents
 - Residents care about location's demographic composition: Bayer, Ferreira & McMillan (2007)

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- New equilibrium prices
 - Need to restore market clearing

This Paper: What are the overall effects of urban renewal programs?

Structural model of residential choice to study welfare and distributional impacts

- Disentangle mechanisms, equilibrium effects, and differences across demographic groups
→ Epple and Sieg (1999), Davis et al. (2019), Tsivanidis (2019), Balboni et al. (2020), Khanna et al. (2020), Couture et al. (2021), Gechter and Tsivanidis (2022)

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New evidence on one of the largest spatially-targeted redevelopment programs in the U.S.

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Focus on PH demolitions in Chicago around the early 2000s

- Affected about 1% of housing stock

Background

Background: Public Housing in Chicago and HOPE VI

Third-largest PH system: 13 high-rise projects and \approx 40k units

Nearly all residents are Black, average HH earnings of \$7k

Serious **infrastructure problems** and **high crime**

Cabrini Green in 2005



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HOPE VI \approx 23,000 demolished units between 1995–2010

Section 8 vouchers to evicted residents

Public Housing Demolitions Late 1990s



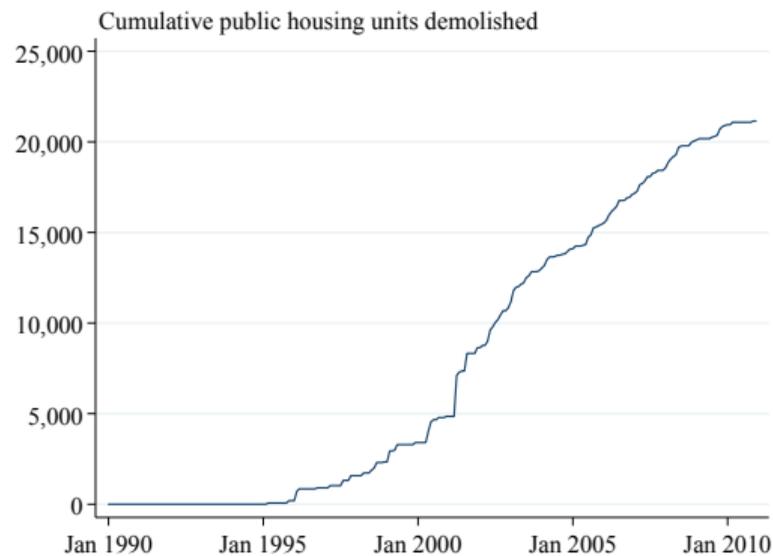
Data and Motivating Facts

Tract-level data:

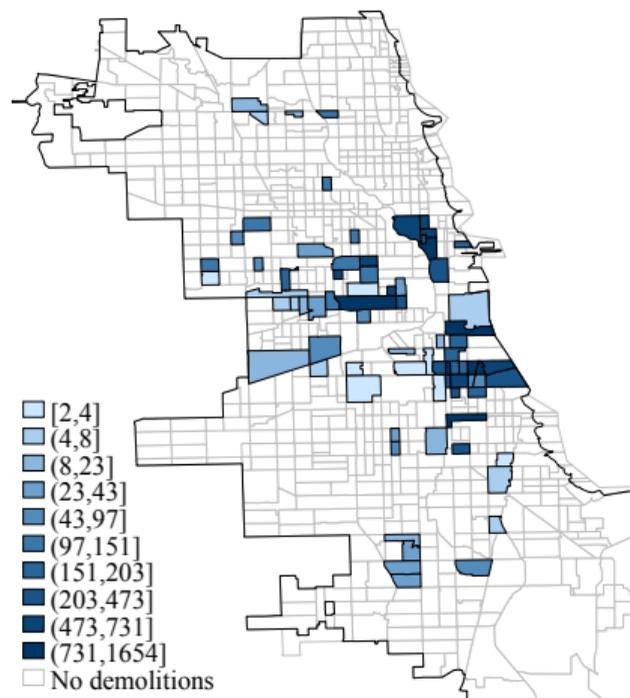
- Chicago Housing Authority: Demolition date and number of units demolished (1995–2010)
- Census Bureau: Population, demographics, and housing values (2000, 2010)

Public Housing Demolitions over Time and Space

Total demolished public housing units



Spatial distribution of public housing demolitions



Tract-level demolitions

Change in Neighborhood Characteristics: Black Population Share and Median House Values

Compare *changes* between 2000–2010 (2016) for tracts with more vs. fewer demolitions

Crime: Aliprantis and Hartley 2015, Sandler 2017, House Prices: Blanco 2021

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Figure 1: Change in Black Population Share

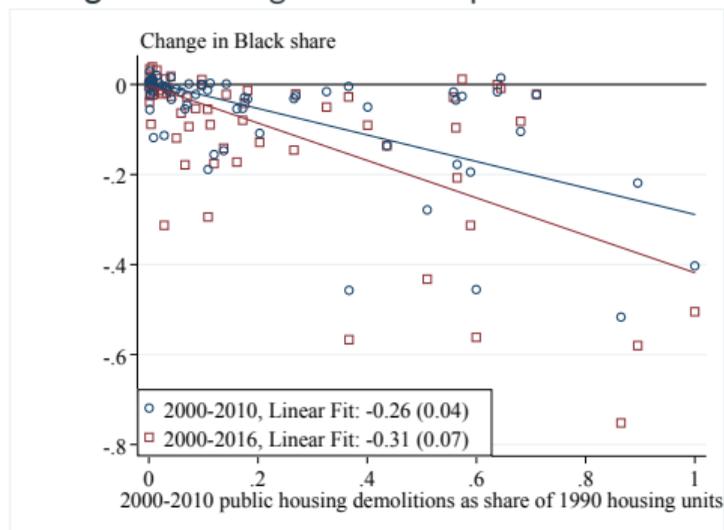
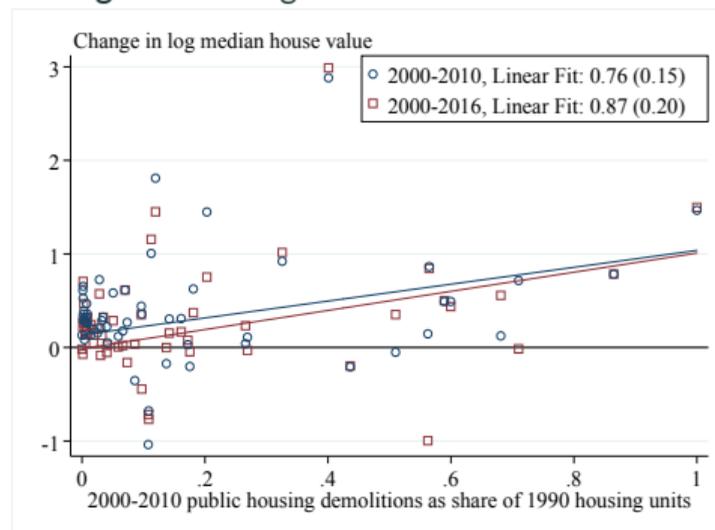


Figure 2: Change in Median House Values



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Housing Demand

Housing Supply

Equilibrium Definition and Model Fit

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Residential Choice Model by Demographic Type

Household of *race* and *income* type k chooses neighborhood j by solving:

$$\max_j V_{jt}^i = \alpha_p^k \ln(p_{jt}) + \alpha_b^k b_{jt} + \alpha_h^k h_{jt} + \alpha_{inc}^k \ln(Inc_{jt}) + \alpha_{PH}^k PH_{jt} + \theta^k x_{jt} + \xi_{jt}^k + \epsilon_{ijt},$$

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- ξ_{jt}^k : unobservable neighborhood characteristics
- ϵ_{ijt} : Type I EV i.i.d. idiosyncratic shock, leading to choice probability:

$$\mathbb{P}_{jt}^k = \frac{\exp(\alpha_p^k \ln(p_{jt}) + \alpha_b^k b_{jt} + \alpha_h^k h_{jt} + \alpha_{Inc}^k \ln(Inc_{jt}) + \alpha_{PH}^k PH_{jt} + \theta^k x_{jt} + \xi_{jt}^k)}{\sum_{j'} \exp(\alpha_p^k \ln(p_{j't}) + \alpha_b^k b_{j't} + \alpha_h^k h_{j't} + \alpha_{Inc}^k \ln(Inc_{j't}) + \alpha_{PH}^k PH_{j't} + \theta^k x_{j't} + \xi_{j't}^k)}$$

Eight demographic types: Race (White, Black, Hispanic, Other) and income groups (Low/High)

Empirical model:

$$\log\left(\frac{\mathbb{P}_{jt}^k}{\mathbb{P}_{0t}^k}\right) = \alpha_p^k \ln(p_{jt}) + \alpha_b^k b_{jt} + \alpha_h^k h_{jt} + \alpha_{inc}^k \ln(Inc_{jt}) + \alpha_{PH}^k PH_{jt} + \theta^k x_{jt} + \xi_{jt}^k \quad (1)$$

→ Construct \mathbb{P}_{jt}^k using share of type k living in tract j in year $t = 2000, 2010$ using census/ACS data

→ Spatially smooth \mathbb{P}_{jt}^k to deal with noise due to small samples and zero shares

Address several threats to identification:

$$\xi_{jt}^k = \lambda_j^k + \lambda_t^k + \tilde{\xi}_{jt}^k$$

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3. Instruments: potential correlation between time-varying characteristics and unobservables $\tilde{\xi}_{jt}^k$

Extend Bayer, Ferreira, and McMillan 2007 to a panel setting

IV details

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BLP instruments (Berry et al. 1995): Exogenous characteristics from tracts > 3 miles away

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- Construct spatial rings from tracts 3-5, 5-10, 10-20 miles

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Intuition

- Consider Hyde Park and Lincoln Park
- Increase in the number of 3-bedroom apartments in Lincoln Park
- Lowers demand for living in Hyde Park
 - Reduction in rents
 - Reduction in high-income HHs that substitute away from Hyde Park

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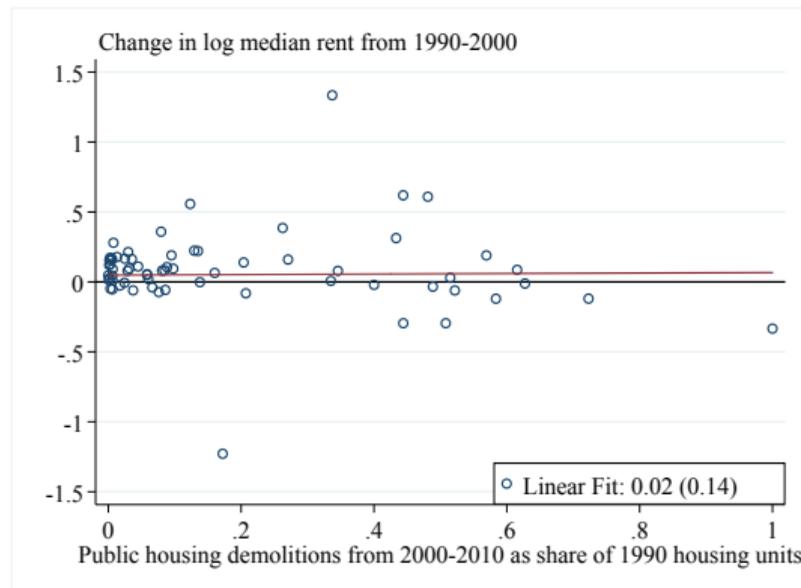
- Consider Hyde Park and Lincoln Park
- Increase in the number of 3-bedroom apartments in Lincoln Park
- Lowers demand for living in Hyde Park
 - Reduction in rents
 - Reduction in high-income HHs that substitute away from Hyde Park
- Requires no correlation with Hyde Park's unobservables

IV details: Three-step approach

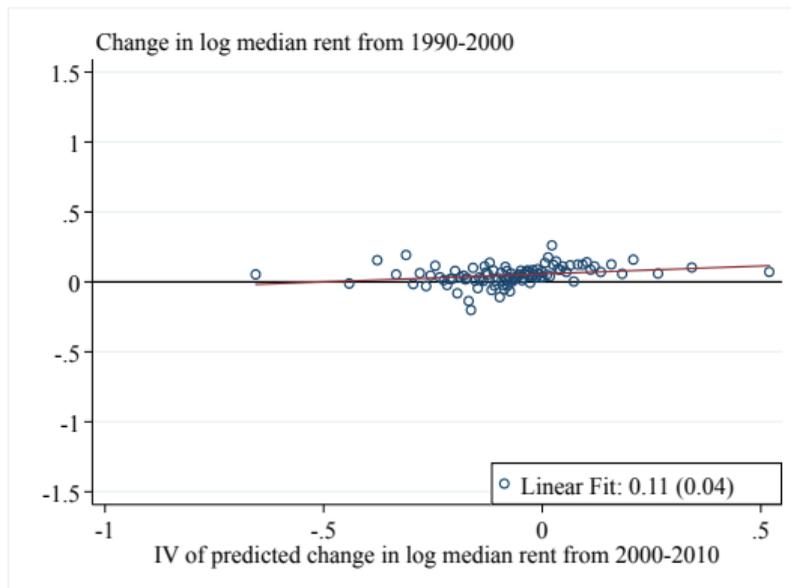
Three-step approach to **increase strength of first-stage** (Davis, Gregory, Hartley, and Tan, 2021):

1. Estimate initial preference parameters using previously discussed instruments
2. Solve for equilibrium rents and demographics with initial estimates and no unobservables $\xi_{jt}^k = 0$
3. Estimate preference parameters adding simulated rent and demographics as instruments

Changes in Log Rents 1990–2010 against PH Demolitions and IV Predicted Log Rent in 2000–2010



(a) Public Housing Demolitions, 2000–2010



(b) IV Predicted Change in Log Rents, 2000–2010

Summary of Preference Estimates

- Negative price elasticities for all groups
→ Low income groups more sensitive to prices

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For example, small average absolute changes in rent coefficient after:

- Controlling for crime ~ 0.001
- Adding more or fewer spatial controls $\sim 0.01-0.02$
- Constructing IVs with more or fewer spatial rings $\sim 0.01-0.02$
- Including region-specific time trends $\sim 0.03-0.05$
- Dropping tracts that are nearby Cabrini-Green ~ 0.02

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Housing supply curve:

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Calibration:

- Housing supply elasticity $\psi = 0.163$ (Baum-Snow and Han 2021 estimates for Chicago)
- Shifter θ_{jt} to match observed equilibrium quantities and prices

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Distinctive equilibrium definition

Requires simultaneous market clearing in

- Housing (prices)
- Demographics

Solve for a $\sim 4,000$ -dimensional fixed point

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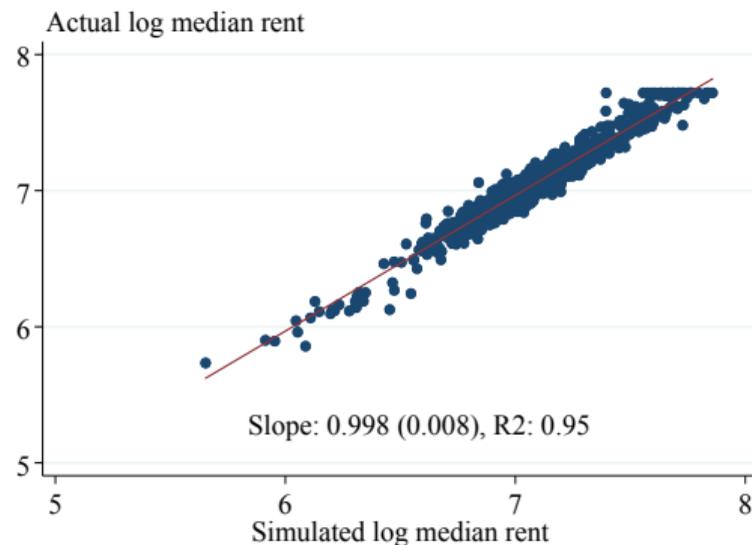
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High Correlation between Actual & Simulated Rent



(a) 2010

Counterfactuals and Welfare

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But effects unevenly distributed across demographic groups

Effects of Demolitions for Non-Hispanic White and Black HHs

Annual Rent Equivalent Under Various Scenarios, White and Black HHs (Median Annual Rent: \$12,288)

Own Demos	Nearby Demos	PH Demographics	Re-sort Demographics	Rent Change	Non-Hisp. White		Black	
					Low Income	High Income	Low Income	High Income
X	X	X	X	X	-	-	-	-

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✓	✓	✓	✓	✓	-\$51	\$56	-\$429	-\$72

Demolitions → welfare gains for high-income white HHs & losses for others (esp. low-income Black)

- More elastic housing supply increases welfare for all groups [More](#)
- Distributing rent income raises welfare gap between LI Black and HI White by \$210 (43%) [More](#)
- “Race-blind” equilibrium → racial preferences account for 40% of welfare changes [More](#)
- Redevelopment in PH areas increases aggregate welfare by \$419 per person [More](#)
- Spatial spillovers and heterogeneity: [More](#)
 - 5.3% average rent increase in treated tracts
 - 1.7% average rent increase in untreated tracts
- The effect of displaced PH households on demand is negligible, both globally and locally.

Conclusion

Empirical evidence of greater inequality by income and race for Chicago PH demolitions:

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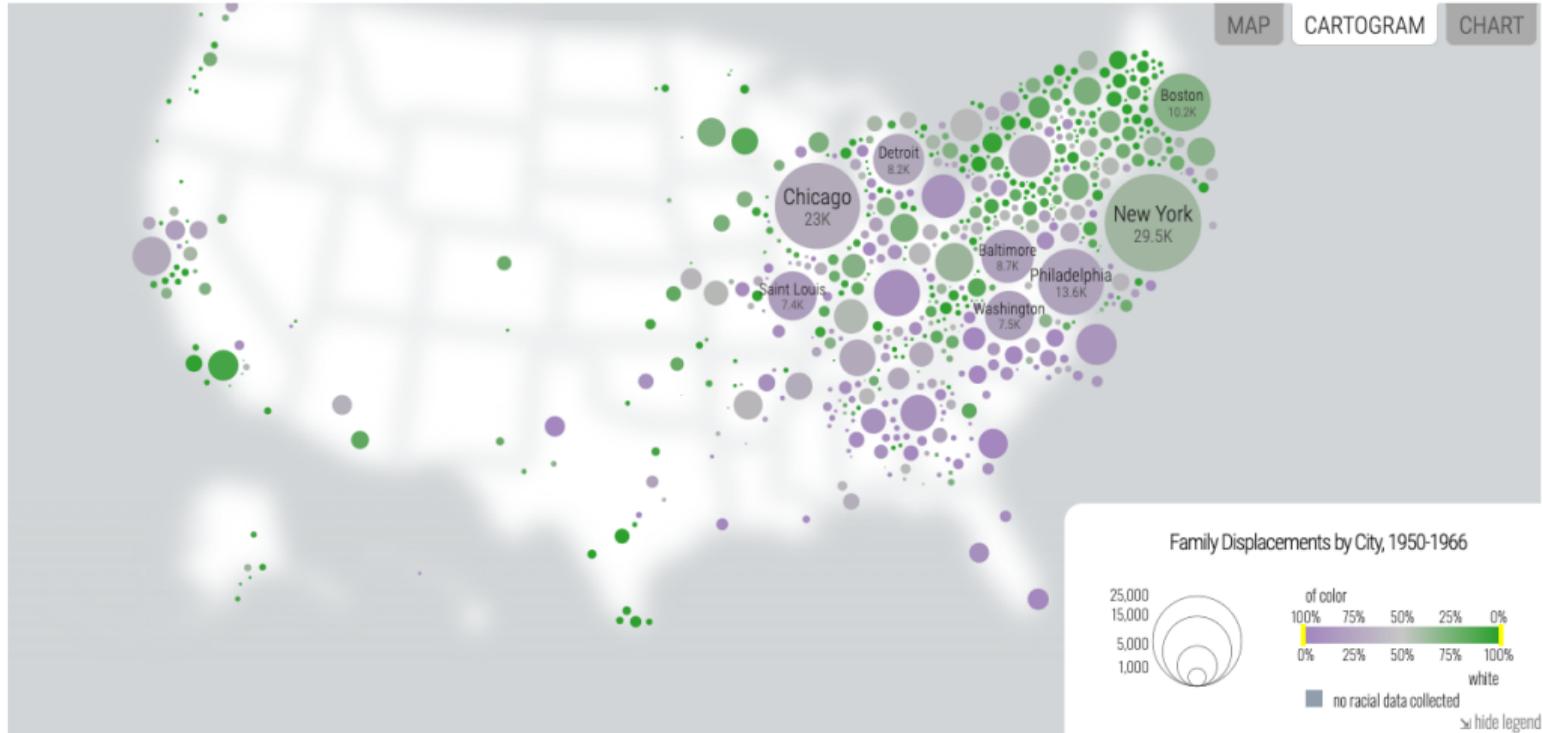
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Increasing the elasticity of housing supply can benefit lower-income minorities in particular

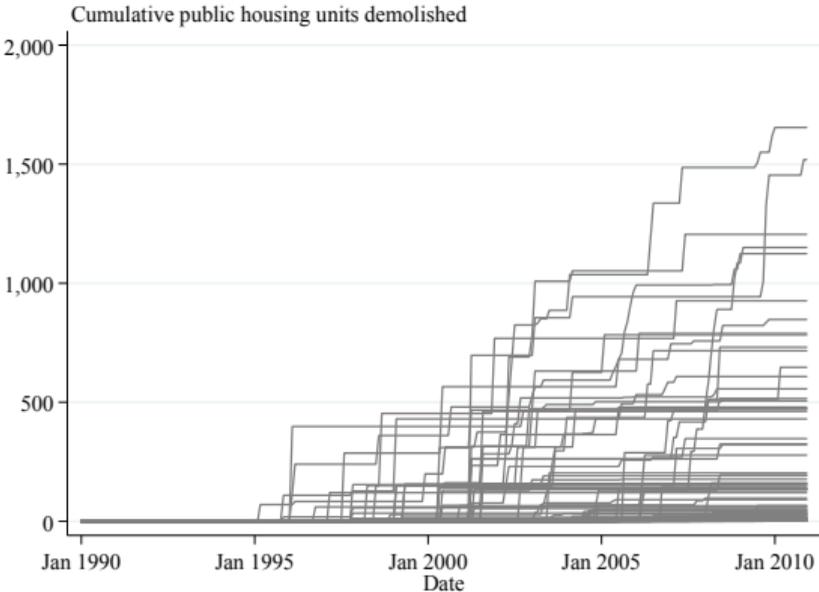
Appendix

Map of forced displacement of Urban Renewal programs 1950-1960



Source: Renewal Inequality, dsl.richmond.edu/panorama/renewal

Cumulative Public Housing Demolitions in Chicago, by Tract



Change in Neighborhood Characteristics: White Population Share and Median HH Income

Figure 5: Change in White Population Share

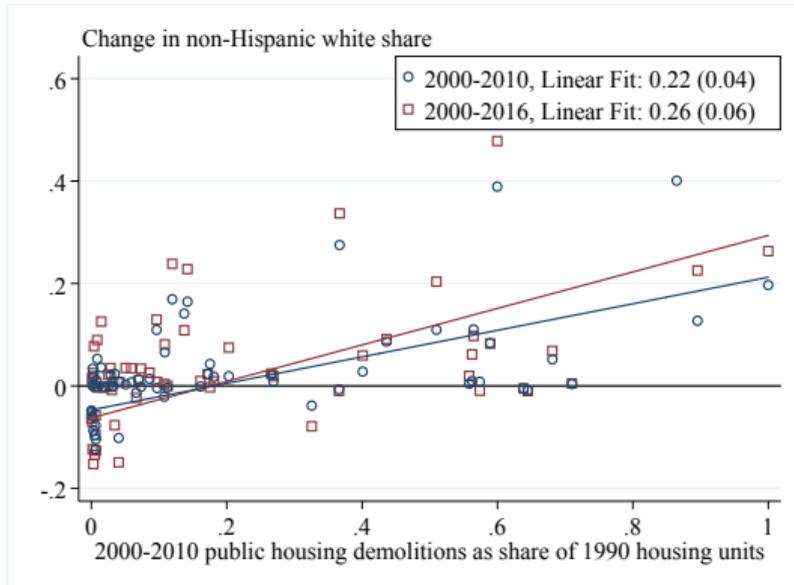
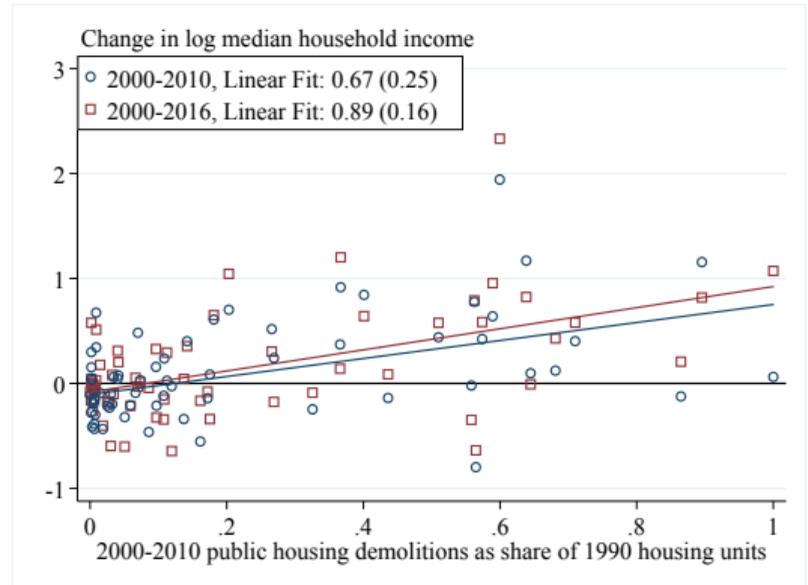


Figure 6: Change in Median Household Income



Preference Estimates for Low Income Groups

IV Estimates of Neighborhood Preference Parameters, Low-Income Only

	Preference parameters for indicated group			
	Non-Hisp White (1)	Black (2)	Hispanic (3)	Other (4)
Log median rent	-0.417*** (0.0486)	-0.168*** (0.0255)	-0.240*** (0.0370)	-0.152*** (0.0296)
Black population share	-0.204*** (0.0787)	0.203*** (0.0406)	0.227*** (0.0537)	-0.158*** (0.0401)
Hispanic population share	-0.0818** (0.0359)	-0.0205 (0.0174)	0.345*** (0.0300)	0.197*** (0.0280)
Log median household income	0.0877*** (0.0211)	0.0132 (0.00933)	0.0259* (0.0152)	0.0631*** (0.0125)
PH units as a share of housing stock	-0.345*** (0.0823)	-0.177*** (0.0472)	-0.255*** (0.0667)	-0.153*** (0.0507)
Observations (tract-by-year)	2,476	2,476	2,476	2,476
Number of tracts	1,238	1,238	1,238	1,238

Preference Estimates for High Income Groups

IV Estimates of Neighborhood Preference Parameters, High-Income Only

	Preference parameters for indicated group			
	Non-Hisp White (1)	Black (2)	Hispanic (3)	Other (4)
Log median rent	-0.0624*** (0.00879)	-0.0520*** (0.0124)	-0.150*** (0.0265)	-0.0265 (0.0169)
Black population share	-0.127*** (0.0147)	0.231*** (0.0218)	0.165*** (0.0346)	-0.107*** (0.0205)
Hispanic population share	-0.133*** (0.00787)	0.0888*** (0.00918)	0.317*** (0.0234)	0.0434** (0.0198)
Log median household income	0.0211*** (0.00395)	0.0130*** (0.00457)	0.0289*** (0.0110)	0.0406*** (0.00694)
PH units as a share of housing stock	-0.0457*** (0.0137)	-0.0839*** (0.0216)	-0.157*** (0.0447)	-0.0294 (0.0298)
Observations (tract-by-year)	2,476	2,476	2,476	2,476
Number of tracts	1,238	1,238	1,238	1,238

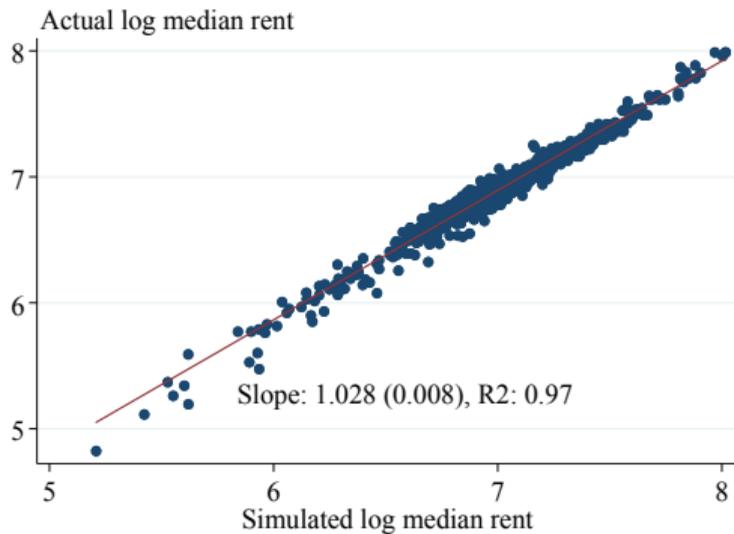
Robustness of Preference Estimates for High-Income White Households

	Baseline (1)	No spatial controls (2)	More spatial controls (3)	Add murder rate (4)	IV rings: 2-3, 3-5 miles (5)	IV rings: 2-3, 3-5, 5-10 miles (6)	Comm. region by year FEs (7)	Comm. division by year FEs (8)	Drop < 1 mile Cabrini-Green (9)
Log rent	-0.0624*** (0.00879)	-0.0672*** (0.00941)	-0.0753*** (0.0101)	-0.0634*** (0.00893)	-0.0768*** (0.0101)	-0.0708*** (0.00956)	-0.0655*** (0.0206)	-0.0717*** (0.0221)	-0.0499*** (0.00832)
Sh. Black	-0.127*** (0.0147)	-0.141*** (0.0170)	-0.112*** (0.0157)	-0.126*** (0.0150)	-0.123*** (0.0180)	-0.115*** (0.0157)	-0.387*** (0.0601)	-0.338*** (0.0707)	-0.125*** (0.0139)
Sh. Hispanic	-0.133*** (0.00787)	-0.130*** (0.00845)	-0.128*** (0.00892)	-0.132*** (0.00790)	-0.140*** (0.00803)	-0.149*** (0.00830)	-0.189*** (0.0203)	-0.243*** (0.0265)	-0.130*** (0.00728)
Log Income	0.0211*** (0.00395)	0.0284*** (0.00463)	0.0225*** (0.00422)	0.0215*** (0.00402)	0.0228*** (0.00463)	0.0219*** (0.00404)	0.0868*** (0.0212)	0.0583* (0.0337)	0.0185*** (0.00337)
PH units	-0.0457*** (0.0137)	-0.0651*** (0.0142)	-0.0629*** (0.0160)	-0.0448*** (0.0139)	-0.0611*** (0.0164)	-0.0575*** (0.0154)	0.00925 (0.0308)	-0.0201 (0.0410)	-0.0355*** (0.0132)

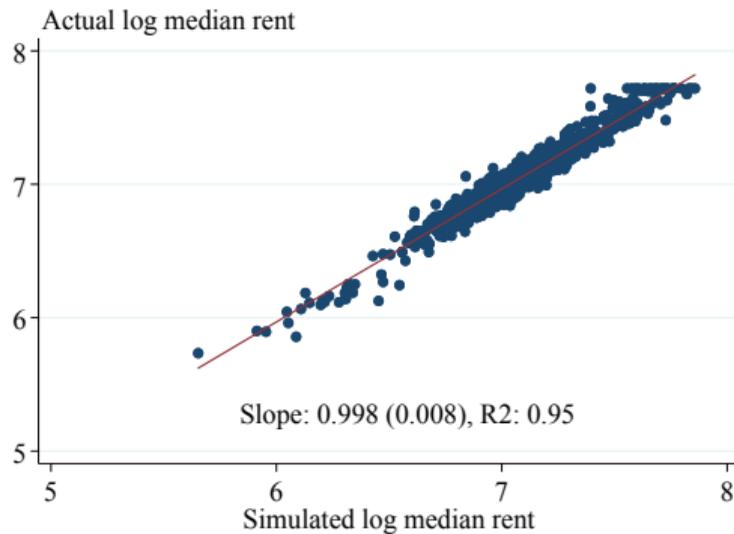
Robustness of Preference Estimates for Low-Income Black Households

	Baseline (1)	No spatial controls (2)	More spatial controls (3)	Add murder rate (4)	IV rings: 2-3, 3-5 miles (5)	IV rings: 2-3, 3-5, 5-10 miles (6)	Comm. region by year FEs (7)	Comm. division by year FEs (8)	Drop < 1 mile Cabrini-Green (9)
Log rent	-0.168*** (0.0255)	-0.197*** (0.0295)	-0.128*** (0.0196)	-0.171*** (0.0258)	-0.178*** (0.0229)	-0.188*** (0.0240)	-0.192** (0.0851)	-0.192** (0.0827)	-0.205*** (0.0294)
Sh. Black	0.203*** (0.0406)	0.247*** (0.0448)	0.158*** (0.0332)	0.204*** (0.0416)	0.157*** (0.0453)	0.186*** (0.0443)	0.908*** (0.169)	0.722*** (0.196)	0.225*** (0.0470)
Sh. Hispanic	-0.0205 (0.0174)	-0.0152 (0.0199)	-0.0227 (0.0144)	-0.0202 (0.0176)	0.0344** (0.0172)	-0.00644 (0.0192)	0.0241 (0.0604)	0.0619 (0.0727)	-0.0268 (0.0198)
Log income	0.0132 (0.00933)	0.0155 (0.0110)	0.0135** (0.00675)	0.0139 (0.00947)	0.0256** (0.0104)	0.0270*** (0.00998)	-0.229*** (0.0614)	-0.147 (0.0913)	0.0143 (0.0108)
PH units	-0.177*** (0.0472)	-0.163*** (0.0526)	-0.124*** (0.0353)	-0.177*** (0.0470)	-0.171*** (0.0454)	-0.188*** (0.0486)	-0.390*** (0.136)	-0.328** (0.145)	-0.199*** (0.0552)

Actual vs. Simulated Log Rent



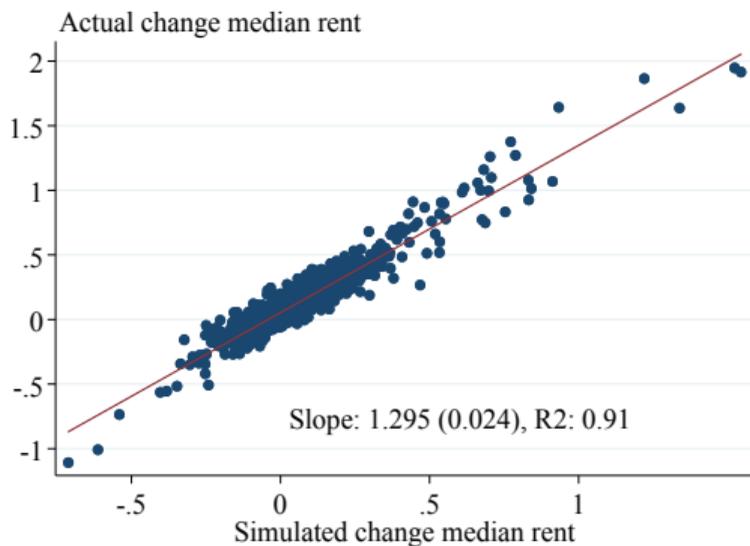
(a) 2000



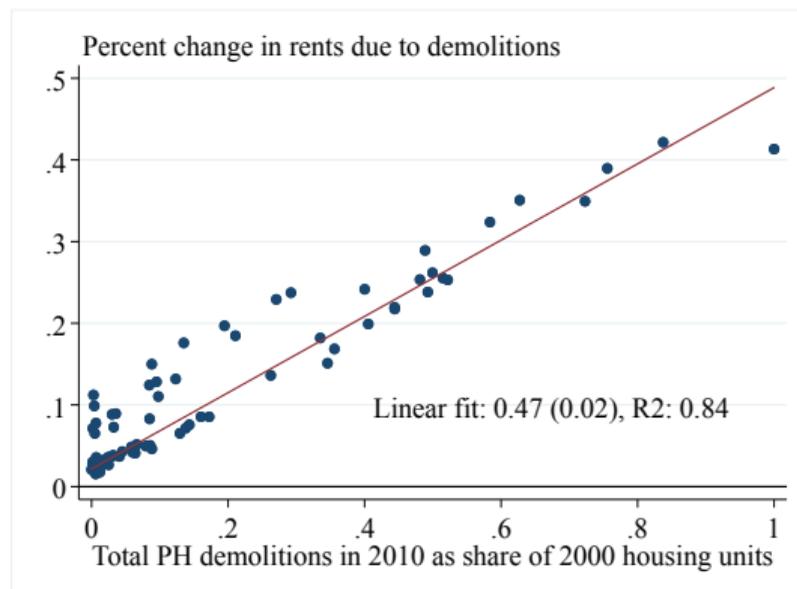
(b) 2010

Model Fit: Rent Changes

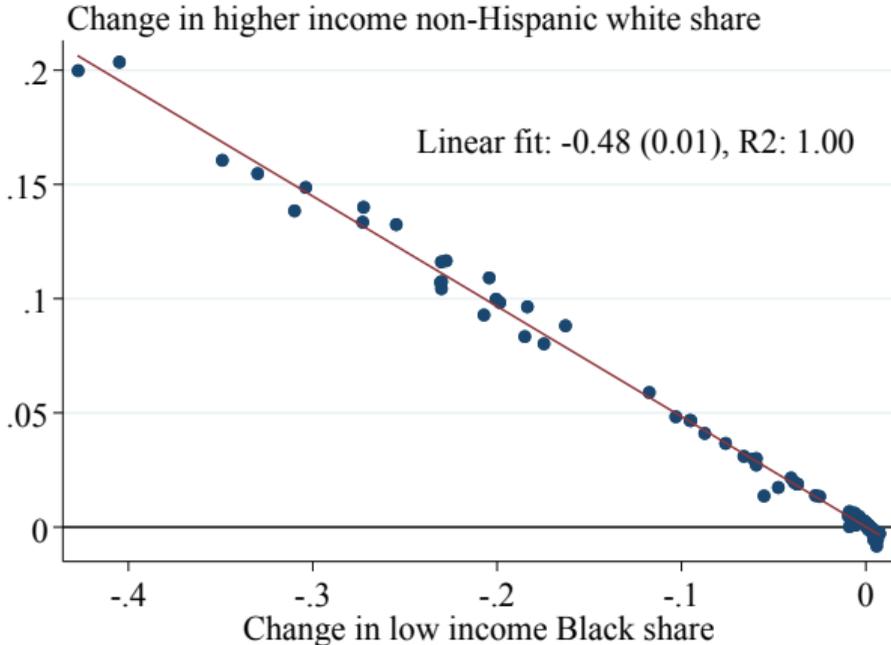
Actual vs. Simulated Log Rent Changes 2000-2010



Rent Changes against Demolitions 2000-2010



Composition effects: High Income White HHs displacing Low Income Black HHs



The Role of Housing Supply

Housing supply elasticity	Change from Baseline (2010 Census)			
	White		Black	
	Low income (1)	High income (2)	Low income (3)	High income (4)
0.000	-313	-241	-1029	-641
0.106	-83	18	-508	-146
0.163 (baseline)	-51	56	-429	-72
0.220	-31	78	-382	-27
0.440	2	115	-300	50

Incorporating Rental Income

	Change from Baseline (2010 Census)			
	White		Black	
	Low income (1)	High income (2)	Low income (3)	High income (4)
Baseline welfare change (no rent redistribution)	-51	56	-429	-72
Average welfare change with rent redistribution	122	329	-366	109
Welfare change for renters	-51	56	-429	-72
Welfare change for homeowners	286	392	-94	264
Homeownership rate	51.3%	81.2%	19.0%	53.9%

“Race-Blind” Counterfactual

	Change from Baseline (2010 Census)			
	White		Black	
	Low income (1)	High income (2)	Low income (3)	High income (4)
Rent equivalent change in welfare from demolitions				
Baseline preferences	-51	56	-429	-72
No race/ethnicity preferences	-30	21	-381	20
Percent change in rent equivalent when eliminating race/ethnicity preferences	41%	-63%	11%	128%

Redevelopment Counterfactual

We replace all PH units with new construction developed by the government

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City-wide prices no longer increase after demolitions, with heterogeneous impacts across locations:

- Tracts with PH with an average rent of \$870 experienced a 28% drop in rents
- Tracts without PH with an average rent of \$1138 experienced a 1% increase in rents

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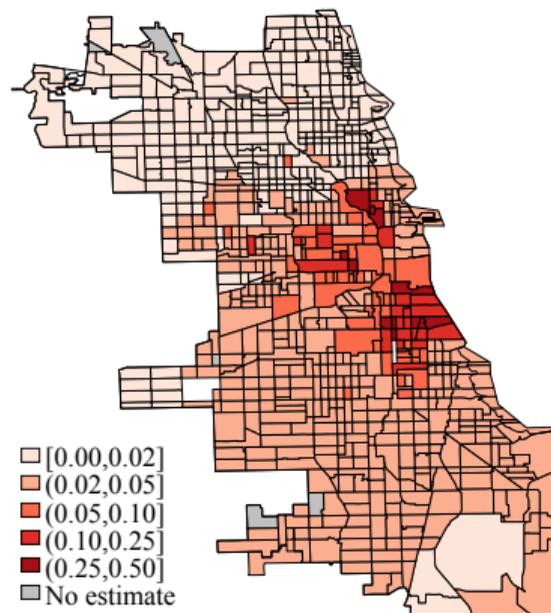
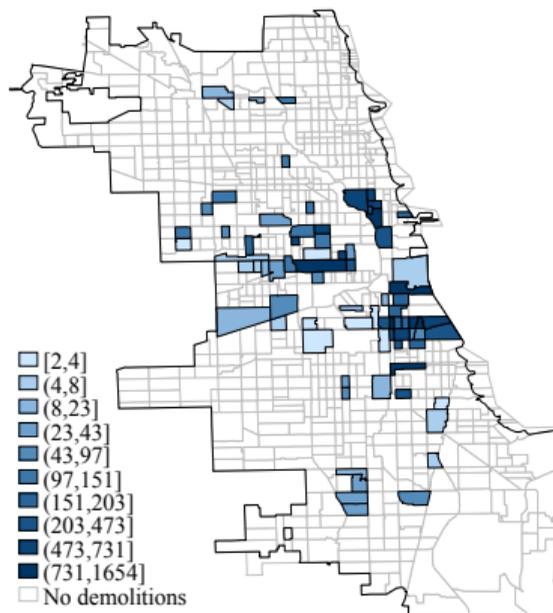
- Tracts with PH with an average rent of \$870 experienced a 28% drop in rents
- Tracts without PH with an average rent of \$1138 experienced a 1% increase in rents

	Change from Baseline (2010 Census)			
	Non-Hispanic White		Black	
	Low income (1)	High income (2)	Low income (3)	High income (4)
Actual redevelopment (baseline)	-51	56	-429	-72
Replace all public housing units with new construction	229	273	90	359

Tract-level Effects of Demolitions on Rents and Spatial Spillovers

Demolitions of PH directly affected 6% of tracts, with spatial spillovers to 27% tracts

→ 2.5% city-wide rent increase



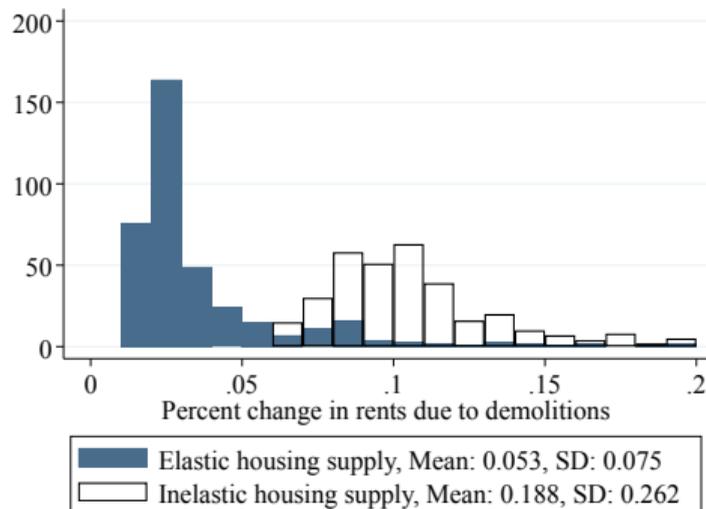
Tract-level Effects of Demolitions on Rents and Spatial Spillovers

Demolitions of PH directly affected 6% of tracts, with spatial spillovers to 27% tracts

→ 2.5% city-wide rent increase

→ 5.3% in directly treated areas, 1.7% in areas more than 1 mile away

Rent Change, Tracts with PH < 1 Mile Away



Rent Change, Tracts with PH > 1 Mile Away

